Sparse Linear Models: Estimation and Approximate Bayesian Inference

Matthias Seeger

Laboratory for Probabilistic Machine Learning Ecole Polytechnique Fédérale de Lausanne

http://lapmal.epfl.ch/







- Denoising
- Natural image statistics
- Wavelet shrinkage
- Image coding

< 47 ▶

Buzzwords



- Denoising
- Natural image statistics
- Wavelet shrinkage
- Image coding

- Feature selection
- ℓ_1 relaxation
- Learning model structure
- Sparse covariance estimation

Buzzwords



- Denoising
- Natural image statistics
- Wavelet shrinkage
- Image coding
- Compressive sensing
- Below the Nyquist limit
- Sparse sampling

- Feature selection
- ℓ_1 relaxation
- Learning model structure
- Sparse covariance estimation



- Denoising
- Natural image statistics
- Wavelet shrinkage
- Image coding
- Compressive sensing
- Below the Nyquist limit
- Sparse sampling

- Feature selection
- ℓ_1 relaxation
- Learning model structure
- Sparse covariance estimation
- Matching/basis pursuit
- Soft/hard thresholding
- {Group, graphical, adaptive} Lasso

Sparsity: A Fundamental Concept

... as simple as possible, but not simpler.

What do you mean with simple?



- All specified elements
- Use each of them a little





Sparsity: A Fundamental Concept

... as simple as possible, but not simpler.

What do you mean with simple?







Sparsity: A Fundamental Concept





Seeger (EPFL)

31/8/2013 4 / 59

- Image modelling
 - Processing
 - Reconstruction
 - Acquisition (sampling)
 - Computational neuroscience





- Image modelling
 - Processing
 - Reconstruction
 - Acquisition (sampling)
 - Computational neuroscience
- Relaxation of combinatorial optimization
 - Maximally sparse reconstruction





- Image modelling
 - Processing
 - Reconstruction
 - Acquisition (sampling)
 - Computational neuroscience
- Relaxation of combinatorial optimization
 - Maximally sparse reconstruction
- Learning dependency structure
 - Meinshausen, Buehlmann
 - Graphical Lasso





- Image modelling
 - Processing
 - Reconstruction
 - Acquisition (sampling)
 - Computational neuroscience
- Relaxation of combinatorial optimization
 - Maximally sparse reconstruction
- Learning dependency structure
 - Meinshausen, Buehlmann
 - Graphical Lasso
- Sparse coding
 - Olshausen, Field
 - Learning image priors





Image Reconstruction





Image Statistics



Whatever images are ...

they are not Gaussian!



Seeger	(EPFL)

- E - N

Bayesian Calibration



U





Computer vision

- Blind deconvolution
- Calibrating camera parameters
- Magnetic resonance imaging
 - Autocalibrating parallel MRI

Bayesian Experimental Design





Seeger ((EPFL)
----------	--------

Outline





- 2 Sparse Estimation
- 3 Sparse Bayesian Inference
- 4 Sparse Estimation vs. Sparse Inference

Outline



Sparse Modelling

- 2 Sparse Estimation
- 3 Sparse Bayesian Inference
- 4 Sparse Estimation vs. Sparse Inference

< 🗇 🕨 < 🖃 >

Sparsity Priors

courtesy Florian Steinke





Best of Both Worlds



$$P(\boldsymbol{u}) \propto \prod_{i=1}^{q} t_i(\boldsymbol{s}_i), \quad \boldsymbol{s} = \boldsymbol{B}\boldsymbol{u}, \quad t_i(\boldsymbol{s}_i) = \boldsymbol{e}^{-\frac{\tau_i}{2}|\boldsymbol{s}_i|^2}$$

Gaussian Prior $P(\mathbf{u})$

- Simple. Fast
- Well understood



< 47 ▶

→ ∃ →

Best of Both Worlds



$$P(\boldsymbol{u}) \propto \prod_{i=1}^{q} t_i(\boldsymbol{s}_i), \quad \boldsymbol{s} = \boldsymbol{B}\boldsymbol{u}, \quad t_i(\boldsymbol{s}_i) = \boldsymbol{e}^{-\tau_i|\boldsymbol{s}_i|}$$

Gaussian Prior $P(\boldsymbol{u})$	Sparsity Prior <i>P</i> (<i>u</i>)	
Simple. FastWell understood	 Better prior for real-world signals (images) 	

Best of Both Worlds



$$P(\boldsymbol{u}) \propto \prod_{i=1}^{q} t_i(\boldsymbol{s}_i), \quad \boldsymbol{s} = \boldsymbol{B}\boldsymbol{u}, \quad t_i(\boldsymbol{s}_i) = \boldsymbol{e}^{-\tau_i|\boldsymbol{s}_i|}$$

Gaussian Prior $P(\boldsymbol{u})$	Sparsity Prior $P(\boldsymbol{u})$	
Simple. FastWell understood	 Better prior for real-world signals (images) 	

Latent Gaussian Representations

- Gaussian scale mixtures
- Super-Gaussian potentials

 $egin{aligned} t(s) &= \int_{\gamma \geq 0} e^{-|s|^2/(2\gamma)} f(\gamma) \, d\gamma \ t(s) &= \max_{\gamma \geq 0} e^{-|s|^2/(2\gamma)} g(\gamma) \end{aligned}$

• We know Gaussian mixtures over means (clustering, EM):

$$P(X) = \sum_{j=1}^{K} \pi_j N(X|\mu_j, \gamma)$$



• We know Gaussian mixtures over means (clustering, EM):

$$P(X) = \sum_{j=1}^{K} \pi_j N(X|\mu_j, \gamma)$$

- What makes t(s) non-Gaussian:
 - More mass close to origin
 - More mass in tails (far from origin)
 - Less mass at moderate distance





• We know Gaussian mixtures over means (clustering, EM):

$$P(X) = \sum_{j=1}^{K} \pi_j N(X|0, \gamma_j)$$

- What makes *t*(*s*) non-Gaussian:
 - More mass close to origin
 - More mass in tails (far from origin)
 - Less mass at moderate distance
 - \Rightarrow Need mixture over scales





$$X = \sqrt{\gamma} Y$$
: $Y \sim N(0, 1), \gamma \sim f(\gamma) I_{\{\gamma \geq 0\}}$

• Many distributions you know:

• Gaussian [:-)].



 $P(X) = N(X|0,\gamma)$



Gaussian Scale Mixtures

$$X = \sqrt{\gamma} Y$$
: $Y \sim N(0, 1), \gamma \sim f(\gamma) I_{\{\gamma \geq 0\}}$

• Many distributions you know:

• Gaussian [:-)]. Spike and slab



$$\boldsymbol{P}(\boldsymbol{X}) = \pi \boldsymbol{N}(\boldsymbol{X}|\boldsymbol{0},\gamma_1) + (1-\pi)\boldsymbol{N}(\boldsymbol{X}|\boldsymbol{0},\gamma_2), \quad \gamma_1 \ll \gamma_2$$





• Many distributions you know:

- Gaussian [:-)]. Spike and slab
- Exponential power ($\alpha \leq$ 2)



$$oldsymbol{P}(oldsymbol{X}) \propto oldsymbol{e}^{- au |oldsymbol{X}|^lpha}, \quad lpha \in (0,2], \; au > oldsymbol{0}$$





• Many distributions you know:

- Gaussian [:-)]. Spike and slab
- Exponential power ($\alpha \leq$ 2)
- Student's t



$$P(X) \propto \left(1+rac{ au}{
u}|X|^2
ight)^{-(
u+1)/2}, \quad au,
u>0$$



$$oldsymbol{X} = \sqrt{\gamma} oldsymbol{Y}$$
: $oldsymbol{Y} \sim oldsymbol{N}(0,1), \, \gamma \sim oldsymbol{f}(\gamma) \mathrm{I}_{\{\gamma \geq oldsymbol{0}\}}$

• Many distributions you know:

- Gaussian [:-)]. Spike and slab
- Exponential power ($\alpha \leq$ 2)
- Student's t



• Duality between P(X) and $f(\gamma)$

West, Biometrika 87

For the Laplace:

$$\frac{\tau}{2}\boldsymbol{e}^{-\tau|\boldsymbol{s}|} = \mathrm{E}[\boldsymbol{N}(\boldsymbol{s}|\boldsymbol{0},\gamma)], \quad \gamma \sim (\tau^2/2)\boldsymbol{e}^{-(\tau^2/2)\gamma}$$



Super-Gaussian Potentials

$$t(s) = \max_{\gamma \ge 0} e^{-|s|^2/(2\gamma)} g(\gamma)$$





イロト イポト イヨト イヨ

Super-Gaussian Potentials

$$t(s) = \max_{\gamma \geq 0} e^{-|s|^2/(2\gamma)} g(\gamma)$$

- t(s) even and positive: Let's look at $|s|^2 \mapsto 2 \log t(s)$
- What's that for a Gaussian $t(s) = N(s|0, \sigma^2)$?





Super-Gaussian Potentials

$$t(s) = \max_{\gamma \geq 0} e^{-|s|^2/(2\gamma)} g(\gamma)$$

- t(s) even and positive: Let's look at $|s|^2 \mapsto 2 \log t(s)$
- What's that for a Gaussian $t(s) = N(s|0, \sigma^2)$? An affine function









Super-Gaussian Potentials

$$t(s) = \max_{\gamma \geq 0} e^{-|s|^2/(2\gamma)} g(\gamma)$$

$$|s|^2 \mapsto 2 \log t(s)$$
 is convex

 Affine → convex: Shift mass to center and tails





Super-Gaussian Potentials

$$t(s) = \max_{\gamma \geq 0} e^{-|s|^2/(2\gamma)} g(\gamma)$$

Sparsity potentials are super-Gaussian

$$|s|^2 \mapsto 2 \log t(s)$$
 is convex

- Affine → convex: Shift mass to center and tails
- Scale mixtures are super-Gaussian









Group Sparsity



$$t_i(s_i) = \max_{\gamma_i \geq 0} e^{-|s_i|^2/(2\gamma_i)} g_i(\gamma_i)$$

• $t_i(s_i)$ depends on absolute value $|s_i|$ only

< 同 > < ∃ >
Group Sparsity



$$t(\boldsymbol{s}_i) = \max_{\gamma_i \geq 0} \underbrace{e^{-\|\boldsymbol{s}_i\|^2/(2\gamma_i)}}_{\propto N(\boldsymbol{s}_i|\boldsymbol{0},\gamma_i\boldsymbol{I})} g_i(\gamma_i)$$

- $t_i(s_i)$ depends on absolute value $|s_i|$ only
- Can just as well plug in vector norm ||s_i||: Nothing but parameter tying

Group Sparsity



$$t(oldsymbol{s}_i) = \max_{\gamma_i \geq 0} e^{-\|oldsymbol{s}_i\|^2/(2\gamma_i)} g_i(\gamma_i)$$

- $t_i(s_i)$ depends on absolute value $|s_i|$ only
- Can just as well plug in vector norm ||s_i||: Nothing but parameter tying
- Useful to structure sparsity: Joint penalization of groups
 ⇒ ℓ₁ − ℓ₂ norms, group Lasso, . . .

Sparse Modelling

Sparsity vs. Super-Gaussianity



Sparse s

•	Ma	ny/r	nos	st s _i	· = 0)	
				•••••			

イロン イ理 とくほとく ほ

Sparse Modelling

Sparsity vs. Super-Gaussianity





A D M A A A M M

Sparse Modelling

Sparsity vs. Super-Gaussianity





- Why call it sparse then?
 - "Super-Gaussian linear model"?
 - Wait until MAP estimation

< ∃ ►

Where Are We?



- Real-world signals are not Gaussian.
 Gaussian assumptions made for convenience
- Super-Gaussian distributions: Trade-off between realistic and tractable
- Latent Gaussian representations:
 - Gaussian scale mixtures
 - Super-Gaussian potentials (max representation)
- Group potentials: Structure your sparsity
- "Sparse" may mean super-Gaussian

Outline







- 3 Sparse Bayesian Inference
- 4 Sparse Estimation vs. Sparse Inference

< 🗇 🕨 < 🖃 >

Image Reconstruction









Maximum a Posteriori (MAP) Estimation

$$u_* = \operatorname{argmax}_{u} P(y|u)P(u)$$

Seener	(EPEL)	
Seeger		

イロト イヨト イヨト イヨト

Sparse Linear Model





< ∃ ►



$$\boldsymbol{u}_{*} = \operatorname{argmin}_{\boldsymbol{u}} \underbrace{\sigma^{-2} \| \boldsymbol{y} - \boldsymbol{X} \boldsymbol{u} \|^{2}}_{-2 \log P(\boldsymbol{y} | \boldsymbol{u})} \underbrace{-2 \sum_{i=1}^{q} \log t_{i}(\boldsymbol{s}_{i})}_{-\log P(\boldsymbol{u})}, \quad \boldsymbol{s} = \boldsymbol{B} \boldsymbol{u}, \ \boldsymbol{y} \in \mathbb{C}^{m}$$

- MAP estimate is sparse.
 - $\boldsymbol{s}_* = \boldsymbol{B} \boldsymbol{u}_*$: No more than *m* nonzero $\boldsymbol{s}_{*,i}$ (iii
 - (if $|s_i| \mapsto -\log t_i(s_i)$ concave)

< 17 ▶







$$\boldsymbol{u}_* = \operatorname{argmin}_{\boldsymbol{u}} \underbrace{\sigma^{-2} \| \boldsymbol{y} - \boldsymbol{X} \boldsymbol{u} \|^2}_{-2 \log P(\boldsymbol{y} | \boldsymbol{u})} \underbrace{-2 \sum_{i=1}^q \log t_i(\boldsymbol{s}_i)}_{-\log P(\boldsymbol{u})}, \quad \boldsymbol{s} = \boldsymbol{B} \boldsymbol{u}, \ \boldsymbol{y} \in \mathbb{C}^m$$

• MAP estimate is sparse.

 $m{s}_* = m{B}m{u}_*$: No more than m nonzero $m{s}_{*,i}$ (if $|s_i| \mapsto -\log t_i(s_i)$ concave)

• MAP convex optimization problem $\Leftrightarrow t_i(s_i)$ log-concave

Sparsity Priors







$$\boldsymbol{u}_{*} = \operatorname{argmin}_{\boldsymbol{u}} \underbrace{\sigma^{-2} \| \boldsymbol{y} - \boldsymbol{X} \boldsymbol{u} \|^{2}}_{-2 \log P(\boldsymbol{y} | \boldsymbol{u})} \underbrace{-2 \sum_{i=1}^{q} \log t_{i}(\boldsymbol{s}_{i})}_{-\log P(\boldsymbol{u})}, \quad \boldsymbol{s} = \boldsymbol{B} \boldsymbol{u}, \ \boldsymbol{y} \in \mathbb{C}^{m}$$

• MAP estimate is sparse.

 $m{s}_* = m{B}m{u}_*$: No more than m nonzero $m{s}_{*,i}$ (if $|s_i| \mapsto -\log t_i(s_i)$ concave)

- MAP convex optimization problem $\Leftrightarrow t_i(s_i)$ log-concave
- Sparse and convex? Laplace potentials (Lasso)

Example: MAP Algorithm



$$\min_{\boldsymbol{u}} \frac{1}{2} \|\boldsymbol{y} - \boldsymbol{X}\boldsymbol{u}\|^2 + \kappa \|\boldsymbol{B}\boldsymbol{u}\|_1$$

• Rewrite: Operator splitting.

< 47 ▶

< ∃ >



$$\min_{\boldsymbol{u},\boldsymbol{s}} \frac{1}{2} \|\boldsymbol{y} - \boldsymbol{X}\boldsymbol{u}\|^2 + \kappa \|\boldsymbol{s}\|_1 \qquad \text{s.t. } \boldsymbol{s} = \boldsymbol{B}\boldsymbol{u}$$

• Rewrite: Operator splitting.

 \Rightarrow Update of each u, s simple (ignoring constraint)



$$\max_{\boldsymbol{b}} \min_{\boldsymbol{u},\boldsymbol{s}} \underbrace{\frac{1}{2} \|\boldsymbol{y} - \boldsymbol{X}\boldsymbol{u}\|^2 + \kappa \|\boldsymbol{s}\|_1 + \lambda \boldsymbol{b}^T (\boldsymbol{B}\boldsymbol{u} - \boldsymbol{s})}_{\text{Lagrangian}}$$

- Rewrite: Operator splitting.
 - \Rightarrow Update of each u, s simple (ignoring constraint)
- Augmented Lagrangian technique (**b** Lagrange multipliers)



$$\max_{\boldsymbol{b}} \min_{\boldsymbol{u},\boldsymbol{s}} \underbrace{\frac{1}{2} \|\boldsymbol{y} - \boldsymbol{X}\boldsymbol{u}\|^2 + \kappa \|\boldsymbol{s}\|_1 + \lambda \boldsymbol{b}^T (\boldsymbol{B}\boldsymbol{u} - \boldsymbol{s}) + \frac{\lambda}{2} \|\boldsymbol{B}\boldsymbol{u} - \boldsymbol{s}\|^2}_{\text{augmented Lagrangian}}$$

- Rewrite: Operator splitting.
 - \Rightarrow Update of each u, s simple (ignoring constraint)
- Augmented Lagrangian technique (**b** Lagrange multipliers)

Example: MAP Algorithm



$\max_{\boldsymbol{b}} \min_{\boldsymbol{u},\boldsymbol{s}} \frac{1}{2} \|\boldsymbol{y} - \boldsymbol{X}\boldsymbol{u}\|^2 + \kappa \|\boldsymbol{s}\|_1 + \lambda \boldsymbol{b}^T (\boldsymbol{B}\boldsymbol{u} - \boldsymbol{s}) + \frac{\lambda}{2} \|\boldsymbol{B}\boldsymbol{u} - \boldsymbol{s}\|^2$

(I) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1)) < ((1))



$$\max_{\boldsymbol{b}} \min_{\boldsymbol{u},\boldsymbol{s}} \frac{1}{2} \|\boldsymbol{y} - \boldsymbol{X}\boldsymbol{u}\|^2 + \kappa \|\boldsymbol{s}\|_1 + \frac{\lambda}{2} \|\boldsymbol{B}\boldsymbol{u} - \boldsymbol{s} + \boldsymbol{b}\|^2 - \frac{\lambda}{2} \|\boldsymbol{b}\|^2$$

Alternating Direction Methods of Multipliers (ADMM)

Iterate:



$$\max_{\boldsymbol{b}} \min_{\boldsymbol{u},\boldsymbol{s}} \frac{1}{2} \|\boldsymbol{y} - \boldsymbol{X}\boldsymbol{u}\|^2 + \kappa \|\boldsymbol{s}\|_1 + \frac{\lambda}{2} \|\boldsymbol{B}\boldsymbol{u} - \boldsymbol{s} + \boldsymbol{b}\|^2 - \frac{\lambda}{2} \|\boldsymbol{b}\|^2$$

Alternating Direction Methods of Multipliers (ADMM)

Iterate:

• Least squares projection (fixed **s**, **b**)

$$\boldsymbol{u} \leftarrow \operatorname{argmin} \frac{1}{2} \| \boldsymbol{y} - \boldsymbol{X} \boldsymbol{u} \|^2 + \frac{\lambda}{2} \| \boldsymbol{B} \boldsymbol{u} - \boldsymbol{s} + \boldsymbol{b} \|^2$$



$$\max_{\boldsymbol{b}} \min_{\boldsymbol{u},\boldsymbol{s}} \frac{1}{2} \|\boldsymbol{y} - \boldsymbol{X}\boldsymbol{u}\|^2 + \kappa \|\boldsymbol{s}\|_1 + \frac{\lambda}{2} \|\boldsymbol{B}\boldsymbol{u} - \boldsymbol{s} + \boldsymbol{b}\|^2 - \frac{\lambda}{2} \|\boldsymbol{b}\|^2$$

Alternating Direction Methods of Multipliers (ADMM)

Iterate:

• Least squares projection (fixed **s**, **b**)

$$\boldsymbol{u} \leftarrow \operatorname{argmin} \frac{1}{2} \| \boldsymbol{y} - \boldsymbol{X} \boldsymbol{u} \|^2 + \frac{\lambda}{2} \| \boldsymbol{B} \boldsymbol{u} - \boldsymbol{s} + \boldsymbol{b} \|^2$$

• Proximal map (fixed **u**, **b**)

$$\boldsymbol{s} \leftarrow \operatorname{argmin} \kappa \|\boldsymbol{s}\|_1 + \frac{\lambda}{2} \|\boldsymbol{B}\boldsymbol{u} - \boldsymbol{s} + \boldsymbol{b}\|^2$$



$$\max_{\boldsymbol{b}} \min_{\boldsymbol{u},\boldsymbol{s}} \frac{1}{2} \|\boldsymbol{y} - \boldsymbol{X}\boldsymbol{u}\|^2 + \kappa \|\boldsymbol{s}\|_1 + \lambda \boldsymbol{b}^T (\boldsymbol{B}\boldsymbol{u} - \boldsymbol{s}) + \frac{\lambda}{2} \|\boldsymbol{B}\boldsymbol{u} - \boldsymbol{s}\|^2$$

Alternating Direction Methods of Multipliers (ADMM)

Iterate:

• Least squares projection (fixed **s**, **b**)

$$\boldsymbol{u} \leftarrow \operatorname{argmin} \frac{1}{2} \| \boldsymbol{y} - \boldsymbol{X} \boldsymbol{u} \|^2 + \frac{\lambda}{2} \| \boldsymbol{B} \boldsymbol{u} - \boldsymbol{s} + \boldsymbol{b} \|^2$$

• Proximal map (fixed **u**, **b**)

$$\boldsymbol{s} \leftarrow \operatorname{argmin} \kappa \|\boldsymbol{s}\|_1 + \frac{\lambda}{2} \|\boldsymbol{B}\boldsymbol{u} - \boldsymbol{s} + \boldsymbol{b}\|^2$$

• Lagrange multiplier update (fixed *u*, *s*)

$$m{b} \leftarrow m{b} + m{B}m{u} - m{s}$$

Example: MRI Reconstruction

$$\min_{\boldsymbol{u}} \frac{1}{2} \|\boldsymbol{y} - \boldsymbol{X}\boldsymbol{u}\|^2 + \kappa \|\boldsymbol{B}\boldsymbol{u}\|_1$$

• $\boldsymbol{X} = \boldsymbol{I}_{J,\cdot} \boldsymbol{F}, \, \boldsymbol{F} \, \text{DFT} \text{ of size } n, \, J \subset \{1, \ldots, n\}$

• Blocks of **B**:

Orthonormal (wavelets), FIR filters (Δ_x , Δ_y)





Example: MRI Reconstruction

$$\min_{\boldsymbol{u}} \frac{1}{2} \|\boldsymbol{y} - \boldsymbol{X}\boldsymbol{u}\|^2 + \kappa \|\boldsymbol{B}\boldsymbol{u}\|_1$$

• $X = I_{J, F}$, F DFT of size $n, J \subset \{1, \ldots, n\}$

 Blocks of **B**: Orthonormal (wavelets), FIR filters (Δ_x, Δ_y)

• Least squares projection:

$$(\boldsymbol{X}^{H}\boldsymbol{X} + \lambda \boldsymbol{B}^{T}\boldsymbol{B})\boldsymbol{u} = \boldsymbol{r} := \boldsymbol{X}^{H}\boldsymbol{y} + \lambda \boldsymbol{B}^{T}(\boldsymbol{s} - \boldsymbol{b})$$







Example: MRI Reconstruction

$$\min_{\boldsymbol{u}} \frac{1}{2} \|\boldsymbol{y} - \boldsymbol{X}\boldsymbol{u}\|^2 + \kappa \|\boldsymbol{B}\boldsymbol{u}\|_1$$

• $\boldsymbol{X} = \boldsymbol{I}_{J,\cdot} \boldsymbol{F}, \, \boldsymbol{F} \, \text{DFT} \text{ of size } n, \, J \subset \{1, \ldots, n\}$

 Blocks of **B**: Orthonormal (wavelets), FIR filters (Δ_x, Δ_y)

• Least squares projection:

$$\left(\boldsymbol{F}^{\mathsf{T}} \boldsymbol{I}_{\cdot,J} \boldsymbol{I}_{J,\cdot} \boldsymbol{F} + \lambda \boldsymbol{B}^{\mathsf{T}} \boldsymbol{B} \right) \boldsymbol{u} = \boldsymbol{r}$$







Example: MRI Reconstruction

$$\min_{\boldsymbol{u}} \frac{1}{2} \|\boldsymbol{y} - \boldsymbol{X}\boldsymbol{u}\|^2 + \kappa \|\boldsymbol{B}\boldsymbol{u}\|_1$$

• $\boldsymbol{X} = \boldsymbol{I}_{J,\cdot} \boldsymbol{F}, \, \boldsymbol{F} \, \text{DFT} \text{ of size } n, \, J \subset \{1, \ldots, n\}$

• Blocks of **B**:

Orthonormal (wavelets), FIR filters (Δ_x , Δ_y)

Least squares projection:

$$F^{T}(I_{\cdot,J}I_{J,\cdot} + \underbrace{\lambda F B^{T} B F^{T}}_{diagonal})Fu = r$$



32/59

31/8/2013





Example: MRI Reconstruction

$$\min_{\boldsymbol{u}} \frac{1}{2} \|\boldsymbol{y} - \boldsymbol{X}\boldsymbol{u}\|^2 + \kappa \|\boldsymbol{B}\boldsymbol{u}\|_1$$

• $\boldsymbol{X} = \boldsymbol{I}_{J,\cdot} \boldsymbol{F}, \, \boldsymbol{F} \, \text{DFT} \text{ of size } n, \, J \subset \{1, \ldots, n\}$

Blocks of **B**:
 Orthonormal (wavelets), FIR filters (Δ_x, Δ_y)

• Least squares projection:

$$(\underline{I_{\cdot,J}I_{J,\cdot}+D})$$
Fu = Fr

 \Rightarrow Two fast Fourier transforms only!





Example: MRI Reconstruction

courtesy Mateusz Malinowski













Seeger (EPFL)

31/8/2013 33 / 59

Example: MRI Reconstruction

courtesy Mateusz Malinowski













Seeger (EPFL)

Sparse Models

Example: MRI Reconstruction

courtesy Mateusz Malinowski













Seeger (EPFL)

31/8/2013

33 / 59

Example: MRI Reconstruction

courtesy Mateusz Malinowski













Seeger (EPFL)

31/8/2013 33 / 59

Where Are We?



• Sparse linear model:

Linear couplings (X, B), super-Gaussian potentials

- MAP estimation:
 - Sparse solution if $|s_i| \mapsto -\log t_i(s_i)$ concave
 - Convex problem if $s_i \mapsto -\log t_i(s_i)$ convex (t_i log-concave)
 - Sparse and convex? Laplace potentials, ℓ_1
- Proximal splitting algorithms: Simple, efficient steps. Parallelizable

Outline



Sparse Modelling

- 2 Sparse Estimation
- Sparse Bayesian Inference
- 4 Sparse Estimation vs. Sparse Inference

4 A N





Maximum a Posteriori (MAP) Estimation

• There are many solutions. Why settle for any single one?

Seeger (EPFL
----------	------

Sparse Models

31/8/2013 36 / 59

A .

→ ∃ →
Integration, not Maximization





Use All Solutions

- Weight each solution by our uncertainty
- Average over them. Integrate, don't maximize

Seeger (EPFL)

Sparse Models

31/8/2013 37 / 59

Robust Model Calibration



$$P(\mathbf{y}|\mathbf{\theta}) = \int P(\mathbf{y}|\mathbf{u},\mathbf{\theta}) P(\mathbf{u}|\mathbf{\theta}) \, d\mathbf{u}$$

Given raw data y, no ground truth u. Calibrate model parameters θ .

- Blind deconvolution (θ blur kernel)
- Multi-frame super-resolution (θ camera parameters, PSF)
- Image coding (θ codebook)
- Learning image priors $(P(\boldsymbol{u}) = P(\boldsymbol{u}|\boldsymbol{\theta}))$

Robust Model Calibration



$$P(\mathbf{y}|\mathbf{\theta}) = \int P(\mathbf{y}|\mathbf{u},\mathbf{\theta})P(\mathbf{u}|\mathbf{\theta}) \, d\mathbf{u}$$

Given raw data y, no ground truth u. Calibrate model parameters θ .

MAP Estimation	Bayesian Inference
 argmax max P(y u)P(u) ^{??} All bets on one θ, all bets on one u, Can work if u much higher-D than θ Additional engineering 	$\operatorname{argmax}_{\theta} \underbrace{\int P(\boldsymbol{y} \boldsymbol{u})P(\boldsymbol{u}) d\boldsymbol{u}}_{\text{likelihood }P(\boldsymbol{y} \theta)}$ • Maximize true likelihood • Account for uncertainty in \boldsymbol{u} : Cues for what θ should be

Bayesian Experimental Design







Variational Bayesian Inference



$$P(\boldsymbol{u}|\boldsymbol{y}) = Z^{-1}P(\boldsymbol{y}|\boldsymbol{u})\prod_{i}t_{i}(s_{i}), \ Z = \int P(\boldsymbol{y}|\boldsymbol{u})\prod_{i}t_{i}(s_{i})\,d\boldsymbol{u}$$

• Bayesian integration over $P(\boldsymbol{u}|\boldsymbol{y})$ intractable

< ∃ >

Variational Bayesian Inference



$$P(\boldsymbol{u}|\boldsymbol{y}) = Z^{-1}P(\boldsymbol{y}|\boldsymbol{u})\prod_{i}t_{i}(s_{i}), \ Z = \int P(\boldsymbol{y}|\boldsymbol{u})\prod_{i}t_{i}(s_{i})\,d\boldsymbol{u}$$

- Bayesian integration over P(u|y) intractable
- Integration tractable for Gaussians Q(u|y)
 ⇒ Approximate P(u|y) by Q(u|y)!

Variational Bayesian Inference



$$P(\boldsymbol{u}|\boldsymbol{y}) = Z^{-1}P(\boldsymbol{y}|\boldsymbol{u})\prod_{i}t_{i}(s_{i}), \ \boldsymbol{Z} = \int P(\boldsymbol{y}|\boldsymbol{u})\prod_{i}t_{i}(s_{i})\,d\boldsymbol{u}$$

- Bayesian integration over P(u|y) intractable
- Integration tractable for Gaussians Q(u|y)
 ⇒ Approximate P(u|y) by Q(u|y)!

Variational approximation

Apply variational principle to fit master function $\log Z$

Super-Gaussian Priors

$$t(s) = \max_{\gamma \ge 0} e^{-\frac{1}{2}(s^2/\gamma + h(\gamma))}$$

$$s^2 \mapsto 2 \log t(s)$$
 is convex

 Affine → convex: Shift mass to center and tails







Super-Gaussian: t(s) even, $\{x = s^2\} \mapsto \{f(x) = 2 \log t(s)\}$ convex.





Super-Gaussian: t(s) even, $\{x = s^2\} \mapsto \{f(x) = 2 \log t(s)\}$ convex.





Super-Gaussian: t(s) even, $\{x = s^2\} \mapsto \{f(x) = 2 \log t(s)\}$ convex.



イロン イ理 とくほとく ほ



Super-Gaussian: t(s) even, $\{x = s^2\} \mapsto \{f(x) = 2 \log t(s)\}$ convex.



イロン イ理 とくほとく ほ



Super-Gaussian: t(s) even, $\{x = s^2\} \mapsto \{f(x) = 2 \log t(s)\}$ convex.



• • • • • • • • • • • •



Super-Gaussian: t(s) even, $\{x = s^2\} \mapsto \{f(x) = 2 \log t(s)\}$ convex.



Super-Gaussian Bounding



$$P(\boldsymbol{u}|\boldsymbol{y}) = rac{P(\boldsymbol{y}|\boldsymbol{u}) imes P(\boldsymbol{u})}{P(\boldsymbol{y})}$$

Sparsity potentials are super-Gaussian

$$t_i(s_i) = \max_{\gamma_i \ge 0} e^{-rac{1}{2}(s_i^2/\gamma_i + h_i(\gamma_i))},$$

 $h(\gamma) := \sum_i h_i(\gamma_i), \ \ \Gamma = \operatorname{diag} \gamma$





Super-Gaussian Bounding



$$P(\boldsymbol{u}|\boldsymbol{y}) = rac{P(\boldsymbol{y}|\boldsymbol{u}) \times P(\boldsymbol{u})}{P(\boldsymbol{y})}$$

Exact representation

$$\log Z = \log \int P(\mathbf{y}|\mathbf{u}) \max_{\gamma} e^{-\frac{1}{2}(\mathbf{s}^{T}\mathbf{\Gamma}^{-1}\mathbf{s} + h(\gamma))} d\mathbf{u}$$



 $t_i(s_i) = \ \max_{\gamma_i \ge 0} e^{-rac{1}{2}(s_i^2/\gamma_i + h_i(\gamma_i))}$

Seeger (EPFL)



Super-Gaussian Bounding



$$P(\boldsymbol{u}|\boldsymbol{y}) = rac{P(\boldsymbol{y}|\boldsymbol{u}) \times P(\boldsymbol{u})}{P(\boldsymbol{y})}$$

Lower bound

$$\log Z$$

$$= \log \int P(\boldsymbol{y}|\boldsymbol{u}) \max_{\boldsymbol{\gamma}} e^{-\frac{1}{2}(\boldsymbol{s}^{T}\boldsymbol{\Gamma}^{-1}\boldsymbol{s}+h(\boldsymbol{\gamma}))} d\boldsymbol{u}$$

$$\geq \max_{\boldsymbol{\gamma}} \log \int P(\boldsymbol{y}|\boldsymbol{u}) e^{-\frac{1}{2}(\boldsymbol{s}^{T}\boldsymbol{\Gamma}^{-1}\boldsymbol{s}+h(\boldsymbol{\gamma}))} d\boldsymbol{u}$$



 $t_i(s_i) = \ \max_{\gamma_i \ge 0} e^{-rac{1}{2}(s_i^2/\gamma_i + h_i(\gamma_i))}$

・日本 ・ヨト・



Super-Gaussian Bounding



$$P(\boldsymbol{u}|\boldsymbol{y}) = rac{P(\boldsymbol{y}|\boldsymbol{u}) \times P(\boldsymbol{u})}{P(\boldsymbol{y})}$$

Lower bound

$$\log Z$$

$$\geq \max_{\gamma} \log \int P(\mathbf{y} | \mathbf{u}) e^{-\frac{1}{2} (\mathbf{s}^{T} \mathbf{\Gamma}^{-1} \mathbf{s} + h(\gamma))} d\mathbf{u}$$

$$= \max_{\gamma} \log Z_{Q}(\gamma) - h(\gamma)/2$$

Gaussian approximation

$$Q(\boldsymbol{u}|\boldsymbol{y}) = Z_{\boldsymbol{Q}}^{-1} P(\boldsymbol{y}|\boldsymbol{u}) e^{-\frac{1}{2}\boldsymbol{s}^{T}\boldsymbol{\Gamma}^{-1}\boldsymbol{s}}, \ \boldsymbol{s} = \boldsymbol{B}\boldsymbol{u}$$



 $t_i(s_i) = \ \max_{\gamma_i \ge 0} e^{-rac{1}{2}(s_i^2/\gamma_i + h_i(\gamma_i))}$



Super-Gaussian Bounding



$$P(\boldsymbol{u}|\boldsymbol{y}) = rac{P(\boldsymbol{y}|\boldsymbol{u}) \times P(\boldsymbol{u})}{P(\boldsymbol{y})}$$

Variational problem: $Q(\boldsymbol{u}|\boldsymbol{y}) \approx P(\boldsymbol{u}|\boldsymbol{y})$

$$\min_{\gamma} \left\{ \phi(\gamma) = -2 \log Z_Q + h(\gamma) \right\}$$

Gaussian approximation

1

$$Q(\boldsymbol{u}|\boldsymbol{y}) = Z_Q^{-1} P(\boldsymbol{y}|\boldsymbol{u}) e^{-\frac{1}{2}\boldsymbol{s}^T \boldsymbol{\Gamma}^{-1} \boldsymbol{s}}, \ \boldsymbol{s} = \boldsymbol{B} \boldsymbol{u},$$
$$Z_Q = \int P(\boldsymbol{y}|\boldsymbol{u}) e^{-\frac{1}{2}\boldsymbol{s}^T \boldsymbol{\Gamma}^{-1} \boldsymbol{s}} d\boldsymbol{u}$$



 $t_i(s_i) = \ \max_{\gamma_i \ge 0} e^{-rac{1}{2}(s_i^2/\gamma_i + h_i(\gamma_i))}$



MAP Estimation and Variational Inference

MAP Estimation

$$\max_{\boldsymbol{u}} \log P(\boldsymbol{u}|\boldsymbol{y})Z$$

$$= \max_{\boldsymbol{u}} \log N(\boldsymbol{y}|\boldsymbol{X}\boldsymbol{u}, \sigma^{2}\boldsymbol{I}) \max_{\boldsymbol{\gamma}} e^{-(s^{T}\Gamma^{-1}\boldsymbol{s}+h(\boldsymbol{\gamma}))/2}$$

$$= \log \int N(\boldsymbol{y}|\boldsymbol{X}\boldsymbol{u}, \sigma^{2}\boldsymbol{I}) \max_{\boldsymbol{\gamma}} e^{-(s^{T}\Gamma^{-1}\boldsymbol{s}+h(\boldsymbol{\gamma}))/2} d\boldsymbol{u}$$

$$= \log \int N(\boldsymbol{y}|\boldsymbol{X}\boldsymbol{u}, \sigma^{2}\boldsymbol{I}) e^{-(s^{T}\Gamma^{-1}\boldsymbol{s}+h(\boldsymbol{\gamma}))/2} d\boldsymbol{u}$$

$$= \log \int N(\boldsymbol{y}|\boldsymbol{X}\boldsymbol{u}, \sigma^{2}\boldsymbol{I}) e^{-(s^{T}\Gamma^{-1}\boldsymbol{s}+h(\boldsymbol{\gamma}))/2} d\boldsymbol{u}$$

$$= \log \int N(\boldsymbol{y}|\boldsymbol{X}\boldsymbol{u}, \sigma^{2}\boldsymbol{I}) e^{-(s^{T}\Gamma^{-1}\boldsymbol{s}+h(\boldsymbol{\gamma}))/2} d\boldsymbol{u}$$

< • > < • > >

- 3 >



Properties of Super-Gaussian Bounding





$$\min_{\gamma} -2\log \int P(\boldsymbol{y}|\boldsymbol{u}) e^{-\frac{1}{2}\boldsymbol{s}^{T}\boldsymbol{\Gamma}^{-1}\boldsymbol{s}} d\boldsymbol{u} + h(\gamma)$$

Super-Gaussian bounding stands out

Seeger, Nickisch, SIAM IS 2011

- Convex problem iff MAP estimation is convex
- Can be solved at much larger scales than others



Properties of Super-Gaussian Bounding





$$\min_{\gamma} -2\log \int P(\boldsymbol{y}|\boldsymbol{u}) e^{-\frac{1}{2}\boldsymbol{s}^{T}\boldsymbol{\Gamma}^{-1}\boldsymbol{s}} d\boldsymbol{u} + h(\gamma)$$

Super-Gaussian bounding stands out

Seeger, Nickisch, SIAM IS 2011

• Convex problem iff MAP estimation is convex

• Can be solved at much larger scales than others

MAP estimation will help solving it!

Seeger ((EPFL)
----------	--------



Towards Scalable Variational Inference

$$\begin{split} \min_{\gamma} -2 \log \int P(\boldsymbol{y} | \boldsymbol{u}) e^{-\frac{1}{2} \boldsymbol{s}^{T} \boldsymbol{\Gamma}^{-1} \boldsymbol{s}} \, d\boldsymbol{u} + h(\gamma) \\ \operatorname{Cov}_{Q}[\boldsymbol{u} | \boldsymbol{y}] &= \boldsymbol{A}^{-1}, \quad \boldsymbol{A} = \sigma^{-2} \boldsymbol{X}^{H} \boldsymbol{X} + \boldsymbol{B}^{T} \boldsymbol{\Gamma}^{-1} \boldsymbol{B} \end{split}$$



• Harder than MAP estimation. But why?



Towards Scalable Variational Inference

$$\begin{split} \min_{\gamma} -2 \log \int P(\boldsymbol{y} | \boldsymbol{u}) e^{-\frac{1}{2} \boldsymbol{s}^{T} \boldsymbol{\Gamma}^{-1} \boldsymbol{s}} \, d\boldsymbol{u} + h(\gamma) \\ \operatorname{Cov}_{Q}[\boldsymbol{u} | \boldsymbol{y}] &= \boldsymbol{A}^{-1}, \quad \boldsymbol{A} = \sigma^{-2} \boldsymbol{X}^{H} \boldsymbol{X} + \boldsymbol{B}^{T} \boldsymbol{\Gamma}^{-1} \boldsymbol{B} \end{split}$$



• Harder than MAP estimation. Because of log |A|.

Super-Gaussian bounding

$$\min_{\boldsymbol{\gamma},\boldsymbol{u}} \left\{ \phi(\boldsymbol{u},\boldsymbol{\gamma}) = \underbrace{\sigma^{-2} \| \boldsymbol{y} - \boldsymbol{X}\boldsymbol{u} \|^2 + \boldsymbol{s}^T \boldsymbol{\Gamma}^{-1} \boldsymbol{s} + h(\boldsymbol{\gamma})}_{\text{MAP criterion}} + \log |\boldsymbol{A}| \right\}$$



Decoupling by Fenchel Duality







• • • • • • • • • • • •

Seeger	(EPFL)
--------	--------

Decoupling by Fenchel Duality



$$\min_{\boldsymbol{\gamma},\boldsymbol{u}_{*}} \phi(\boldsymbol{u}_{*},\boldsymbol{\gamma}) = \min_{\boldsymbol{\gamma},\boldsymbol{u}_{*}} \underbrace{\log |\boldsymbol{A}(\boldsymbol{\gamma}^{-1})|}_{\text{concave}} + \underbrace{\phi_{\cup}(\boldsymbol{u}_{*},\boldsymbol{\gamma})}_{\text{convex}}$$

Fenchel duality

$$\log |\boldsymbol{A}(\boldsymbol{\gamma}^{-1})| = \min_{\boldsymbol{z}} \boldsymbol{z}^{T}(\boldsymbol{\gamma}^{-1}) - \boldsymbol{g}^{*}(\boldsymbol{z})$$



Decoupling by Fenchel Duality



$$\log |\boldsymbol{A}(\gamma^{-1})| + \phi_{\cup}(\boldsymbol{u}_{*}, \gamma) = \min_{\boldsymbol{z}} \underbrace{\boldsymbol{z}^{T}(\gamma^{-1}) + \phi_{\cup}(\boldsymbol{u}_{*}, \gamma) - \boldsymbol{g}^{*}(\boldsymbol{z})}_{\phi_{\boldsymbol{z}}(\boldsymbol{u}_{*}, \gamma) \text{ (convex, decoupled)}}$$

Fenchel duality

$$\log |\boldsymbol{A}(\boldsymbol{\gamma}^{-1})| = \min_{\boldsymbol{z}} \boldsymbol{z}^{T}(\boldsymbol{\gamma}^{-1}) - \boldsymbol{g}^{*}(\boldsymbol{z})$$



< 47 ▶



Double loop algorithm

Seeger et.al., NIPS 2009; insp. by Wipf et.al., NIPS 2008

< 47 ▶

- 3 >

• Inner loop optimization: $\min_{\gamma} \min_{\boldsymbol{u}_*} \phi_{\boldsymbol{z}}(\boldsymbol{u}_*, \gamma) + g^*(\boldsymbol{z})$ [fixed \boldsymbol{z}]

$$\min_{\boldsymbol{u}_*} \min_{\boldsymbol{\gamma}} \sigma^{-2} \|\boldsymbol{y} - \boldsymbol{X} \boldsymbol{u}_*\|^2 + \boldsymbol{z}^T (\boldsymbol{\gamma}^{-1}) + \boldsymbol{s}_*^T \boldsymbol{\Gamma}^{-1} \boldsymbol{s}_* + \boldsymbol{h}(\boldsymbol{\gamma})$$



Double loop algorithm

Seeger et.al., NIPS 2009; insp. by Wipf et.al., NIPS 2008

• Inner loop optimization: $\min_{\gamma} \min_{\boldsymbol{u}_*} \phi_{\boldsymbol{z}}(\boldsymbol{u}_*, \gamma) + g^*(\boldsymbol{z})$ [fixed \boldsymbol{z}] Smoothed MAP Reconstruction

$$\min_{\bm{u}_*} \sigma^{-2} \|\bm{y} - \bm{X} \bm{u}_*\|^2 - 2\sum_{i=1}^q \log t_i \left(\sqrt{z_i + s_{*i}^2} \right), \quad z_i > 0$$





Tangent :
$$\boldsymbol{z} \leftarrow \nabla_{\boldsymbol{\gamma}^{-1}} \log |\boldsymbol{A}|, \quad \boldsymbol{A} = \sigma^{-2} \boldsymbol{X}^{H} \boldsymbol{X} + \boldsymbol{B}^{T} \boldsymbol{\Gamma}^{-1} \boldsymbol{B}$$





$$\boldsymbol{z} \leftarrow \nabla_{\boldsymbol{\gamma}^{-1}} \log |\boldsymbol{A}| = \operatorname{diag}(\boldsymbol{B}\boldsymbol{A}^{-1}\boldsymbol{B}^T) = (\operatorname{Var}_Q[\boldsymbol{s}_i|\boldsymbol{y}])$$

Reductions



Computational primitives driving large scale inference

• Penalized least squares (\approx MAP estimation)

$$\min_{\bm{u}_{*}} \sigma^{-2} \|\bm{y} - \bm{X} \bm{u}_{*}\|^{2} - 2 \sum_{i=1}^{q} \log t_{i} \left(\sqrt{z_{i} + s_{*i}^{2}} \right)$$

- MAP special case: $z_i = 0$
- Scalable algorithms (thanks to MAP "gold rush")

Reductions



Computational primitives driving large scale inference

• Penalized least squares (\approx MAP estimation)

$$\min_{\boldsymbol{u}_*} \sigma^{-2} \|\boldsymbol{y} - \boldsymbol{X} \boldsymbol{u}_*\|^2 - 2 \sum_{i=1}^q \log t_i \left(\sqrt{z_i + s_{*i}^2} \right)$$

• MAP special case: $z_i = 0$

Scalable algorithms (thanks to MAP "gold rush")

Gaussian variances

diag⁻¹(
$$\boldsymbol{B}\boldsymbol{A}^{-1}\boldsymbol{B}^{T}$$
), $\boldsymbol{A} = \sigma^{-2}\boldsymbol{X}^{H}\boldsymbol{X} + \boldsymbol{B}^{T}\boldsymbol{\Gamma}^{-1}\boldsymbol{B}$

- More difficult
- Methods from numerical mathematics, spatial statistics

Where Are We?



- Bayesian inference: Optimization over distributions. Variational approximations: Relaxations thereof
- Super-Gaussian bounding: Exploit latent Gaussian representations of t_i
 - Convex iff MAP estimation is convex
 - Scalable by reductions (double loop algorithm)

Where Are We?



- Bayesian inference: Optimization over distributions.
 Variational approximations: Relaxations thereof
- Super-Gaussian bounding: Exploit latent Gaussian representations of t_i
 - Convex iff MAP estimation is convex
 - Scalable by reductions (double loop algorithm)
- Other relaxations available
- General variational inference
- Randomized approximations (MCMC, brief Gibbs sampling)

Seeger, J. Phys. Conf. 2009 Nickisch et.al., JMLR 2008

Wainwright, Jordan, FTML 2008

Teh et.al., JMLR 2003 Roth, Black, IJCV 2009

Outline



Sparse Modelling

- 2 Sparse Estimation
- 3 Sparse Bayesian Inference
- 4 Sparse Estimation vs. Sparse Inference




- When do I get exact zeros?
- Why is sparse inference more expensive than sparse estimation?
- Can I drive Bayesian experimental design with sparse estimation (RVM/ARD)?

Automatic Relevance Determination Tipping 2001, Wipf et.al. 2004



$$\min_{\gamma} \left\{ \phi_{\mathsf{ARD}}(\gamma) = -2 \log \int P(\boldsymbol{y}|\boldsymbol{u}) N(\boldsymbol{s}|\boldsymbol{0},\boldsymbol{\Gamma}) \, d\boldsymbol{u} \right\}, \quad \boldsymbol{s} = \boldsymbol{B} \boldsymbol{u}$$

- ARD ↔ Relevance Vector Machine
- Sparsity by $\gamma_i \rightarrow 0$
- Sparse estimation, not sparse inference: Seeger, Wipf, IEEE SPM 2010 Zero-temperature limit of variational inference with Student t prior
- Algorithms:
 - Sequential greedy
 - Double loop (reweighted ℓ_1)

Tipping, Faul, AISTATS 2003 Wipf *et.al.*, NIPS 2008

Exact Sparsity Kills Posterior Uncertainty



$$\gamma_i = \mathbf{0} \quad \Rightarrow \quad \mathrm{E}_{\boldsymbol{Q}}[\boldsymbol{s}_i^2 | \boldsymbol{y}] = \mathbf{0}$$

• Exact sparsity controlled by γ_i

Exact Sparsity Kills Posterior Uncertainty

$$\gamma_i = 0 \quad \Rightarrow \quad \mathrm{E}_Q[s_i^2|\mathbf{y}] = 0$$

- Exact sparsity controlled by γ_i
- Exact sparsity kills posterior uncertainty:

$$\boldsymbol{\gamma}_J = \mathbf{0} \quad \Rightarrow \quad \mathrm{E}_{\boldsymbol{Q}}[\boldsymbol{s}_J | \boldsymbol{y}] = \mathbf{0}, \ \mathrm{Cov}_{\boldsymbol{Q}}[\boldsymbol{s}_J | \boldsymbol{y}] = \mathbf{0}$$



Exact Sparsity Kills Posterior Uncertainty

$$\gamma_i = \mathbf{0} \quad \Rightarrow \quad \mathrm{E}_{\boldsymbol{Q}}[\boldsymbol{s}_i^2 | \boldsymbol{y}] = \mathbf{0}$$

- Exact sparsity controlled by γ_i
- Exact sparsity kills posterior uncertainty:

$$\gamma_J = \mathbf{0} \quad \Rightarrow \quad \mathrm{E}_Q[\mathbf{s}_J | \mathbf{y}] = \mathbf{0}, \ \mathrm{Cov}_Q[\mathbf{s}_J | \mathbf{y}] = \mathbf{0}$$

• Good for computation: Heavy scaling only in nonzeros $\|\gamma\|_0$



Exact Sparsity Kills Posterior Uncertainty

$$\gamma_i = 0 \quad \Rightarrow \quad \operatorname{E}_{\boldsymbol{Q}}[\boldsymbol{s}_i^2 | \boldsymbol{y}] = 0$$

- Exact sparsity controlled by γ_i
- Exact sparsity kills posterior uncertainty:

$$\gamma_J = \mathbf{0} \quad \Rightarrow \quad \mathrm{E}_Q[\mathbf{s}_J | \mathbf{y}] = \mathbf{0}, \ \operatorname{Cov}_Q[\mathbf{s}_J | \mathbf{y}] = \mathbf{0}$$

- Good for computation: Heavy scaling only in nonzeros $\|\gamma\|_0$
- Bad for Bayesian inference: Uncertainty is eliminated (the more so, the less data!)
 - Bayesian experimental design: Ji, Carin, ICML 2007 Cannot be based on sparse estimation



Sparse Estimation vs. Sparse Inference

Sparse Estimation

Sparse Inference

э



Sparse Estimation vs. Sparse Inference

(EPA) ÉCOLE POLYTECHNIQU FÉDÉRALE DE LAUSANN

Sparse Estimation

Sparse Inference

• Encourages $\gamma_i \rightarrow 0$

• Forbids $\gamma_i \rightarrow 0$ $(\phi_{\text{SGB}} \rightarrow \infty)$





Seeger, Nickisch, SIAM IS 2001



When do I get exact zeros?
 Controlled by γ_i → 0.
 Happens for sparse estimation, not for sparse inference





- When do I get exact zeros?
 Controlled by γ_i → 0.
 Happens for sparse estimation, not for sparse inference
- Why is sparse inference more expensive than sparse estimation? Because sparse inference maintains posterior uncertainty (full covariance)





- When do I get exact zeros?
 Controlled by γ_i → 0.
 Happens for sparse estimation, not for sparse inference
- Why is sparse inference more expensive than sparse estimation? Because sparse inference maintains posterior uncertainty (full covariance)
- Can I drive Bayesian experimental design with sparse estimation? No free lunch! Sparse estimation kills posterior uncertainty (highly degenerate covariance)

• Variational Bayesian inference very active field

- Loopy belief propagation and generalizations
- Convex relaxations. LP relaxations
- Gaussian/discrete Markov random fields

Wainwright, Jordan FTML 2008

Variational Bayesian inference very active field

- Loopy belief propagation and generalizations
- Convex relaxations. LP relaxations
- Gaussian/discrete Markov random fields
- Broad application impact
 - Coding, information transmission
 - Expert systems
 - · Low level computer vision, adaptive robotics and control
 - Discrete optimization
 - Geostatistics, spatial modelling

Wainwright, Jordan FTML 2008

Mézard, Montanari, 2009

∃ >

Sparse inference beyond MAP estimation

- Robust reconstruction
- Active, adaptive data acquisition
- Learning for inverse problems
- Sequential decision-making

Sparse inference beyond MAP estimation

- Robust reconstruction
- Active, adaptive data acquisition
- Learning for inverse problems
- Sequential decision-making
- Bayesian experimental design, Bayesian optimization
 - Medical imaging sampling optimization
 - Computational photography
 - Intelligent user interfaces
 - Active calibration of cameras

Sparse inference beyond MAP estimation

- Robust reconstruction
- Active, adaptive data acquisition
- Learning for inverse problems
- Sequential decision-making
- Bayesian experimental design, Bayesian optimization
 - Medical imaging sampling optimization
 - Computational photography
 - Intelligent user interfaces
 - Active calibration of cameras
- Modern variational inference algorithms: Layers on top of what you already know
 - Penalized least squares (MAP) reconstruction
 - Gaussian covariance approximation (PCA)

Software and Acknowledgments

glm-ie: Toolbox by Hannes Nickisch

mloss.org/software/view/269/

- Generalized sparse linear models
- MAP reconstruction and variational Bayesian inference (double loop algorithm for super-Gaussian bounding)

Hannes Nickisch (Philips Hamburg, ex-MPI Tübingen)
Rolf Pohmann, Bernhard Schölkopf (MPI Tübingen)

Matlab 7.x, GNU Octave 3.2.x

David Wipf (MSR China)



